

# Stochastic Analysis

## Exercise Sheet 1

November 3/November 10, 2011

- 1 Facts about the normal distribution
- 2 Stopping times and the optional stopping theorem
- 3 Additional remarks

## Theorem

*For a random variable  $X \sim \mathcal{N}(\mu, \sigma^2)$ ,  $a \in \mathbb{R}$  and  $b \in \mathbb{R}$  we have*

$$aX + b \sim \mathcal{N}(a\mu + b, a^2\sigma^2).$$

## Remark

*Here we allow the degenerate case of a constant and we say  $c \sim \mathcal{N}(c, 0)$  for every  $c \in \mathbb{R}$ .*

### Theorem

*For two independent random variables  $X \sim \mathcal{N}(\mu, \sigma^2)$  and  $Y \sim \mathcal{N}(\nu, \rho^2)$  we have*

$$X + Y \sim \mathcal{N}(\mu + \nu, \sigma^2 + \rho^2).$$

## Definition

A random vector  $\vec{X} := (X_1, X_2, \dots, X_n)^{tr}$  has a (multivariate) normal distribution iff

$$\langle (X_1, X_2, \dots, X_n), (a_1, a_2, \dots, a_n) \rangle = \sum_{i=1}^n a_i X_i$$

has a one-dimensional normal distribution for all  $a_1, a_2, \dots, a_n \in \mathbb{R}$ . We write  $\vec{X} \sim \mathcal{N}(\vec{\mu}, \Sigma)$  where

$$\vec{\mu} := (E[X_1], E[X_2], \dots, E[X_n])^{tr}$$

is the mean vector of  $\vec{X}$  and

$$\Sigma := \begin{pmatrix} \text{Cov}(X_1, X_1) & \text{Cov}(X_1, X_2) & \dots & \text{Cov}(X_1, X_n) \\ \text{Cov}(X_2, X_1) & \text{Cov}(X_2, X_2) & \dots & \dots \\ \dots & \dots & \dots & \dots \\ \text{Cov}(X_n, X_1) & \dots & \dots & \text{Cov}(X_n, X_n) \end{pmatrix}$$

## Remark:

The characteristic function can also be defined for a vector  $\vec{X} := (X_1, X_2, \dots, X_n)^{tr}$  of random variables and is given by

$$\varphi_{\vec{X}} : \mathbb{R}^n \rightarrow \mathbb{C}$$

$$\vec{t} = (t_1, t_2, \dots, t_n)^{tr} \mapsto \mathbb{E} \left[ \exp \left( i \langle \vec{t}, \vec{X} \rangle \right) \right] = \mathbb{E} \left[ \exp \left( i \sum_{j=1}^n t_j X_j \right) \right].$$

One can show that the characteristic function of a random vector  $\vec{X} := (X_1, X_2, \dots, X_n)^{tr}$  with a  $\mathcal{N}(\vec{\mu}, \Sigma)$  distribution is given by

$$\varphi_{\vec{X}}(\vec{t}) = \exp \left( i \cdot \vec{t}^{tr} \cdot \vec{\mu} - \frac{1}{2} \vec{t}^{tr} \cdot \Sigma \cdot \vec{t} \right).$$

Since the distribution of every random vector is uniquely determined by its characteristic function it turns out that the normal distribution  $\mathcal{N}(\vec{\mu}, \Sigma)$  is completely characterized by  $\vec{\mu}$  and  $\Sigma$ .

## Theorem

Consider a random vector

$$\vec{X} := \begin{pmatrix} X_1 \\ X_2 \\ \vdots \\ X_n \end{pmatrix}$$

with a  $\mathcal{N}(\vec{\mu}, \Sigma)$  distribution, some matrix  $A \in \mathbb{R}^{k \times n}$  and some vector  $\vec{b} \in \mathbb{R}^k$ . Then

$$A\vec{X} + \vec{b}$$

has a  $\mathcal{N}(A\vec{\mu} + \vec{b}, A\Sigma A^{tr})$ -distribution.

**Proof:** We will show this Theorem in Exercise 2.1!

## Theorem

Consider a random vector  $\vec{X} := (X_1, X_2, \dots, X_n)^{tr}$  where  $\vec{X} \sim \mathcal{N}(\vec{\mu}, \Sigma)$ . If  $\Sigma$  is not singular, i.e.  $\det(\Sigma) > 0$  then the inverse  $\Sigma^{-1}$  exists and  $\vec{X}$  has the density

$$f_{\vec{X}}(\vec{x}) = \left(\frac{1}{2\pi}\right)^{\frac{n}{2}} \cdot \frac{1}{\sqrt{\det(\Sigma)}} \cdot \exp\left(-\frac{1}{2}(\vec{x} - \vec{\mu})^{tr} \cdot \Sigma^{-1} \cdot (\vec{x} - \vec{\mu})\right)$$

with respect to the  $n$ -dimensional Lebesgue-measure.



### Theorem

*The components of a vector  $\vec{X}$  with a normal distribution are independent iff they are uncorrelated.*

### Corollary

*The components of a vector  $\vec{X}$  with a normal distribution are independent iff they are pairwise independent.*

## Proof:

We only have to show that uncorrelatedness implies independence: By assumption we have  $\text{Cov}(X_i, X_j) = 0$  for  $i \neq j$ . Therefore the covariance matrix is diagonal with entries  $\sigma_1^2, \sigma_2^2, \dots, \sigma_n^2$  which denote the variance of  $X_1, X_2, \dots, X_n$ . If one of the variances is 0 then that component is degenerate (i.e. constant with probability 1) and hence it is independent of all other components. We can therefore assume w.l.o.g. that we have  $\sigma_1^2, \sigma_2^2, \dots, \sigma_n^2 > 0$  which implies that the inverse  $\Sigma^{-1}$  exists and is also diagonal with entries  $1/\sigma_1^2, 1/\sigma_2^2, \dots, 1/\sigma_n^2$ . Therefore the density function of  $\vec{X}$  simplifies into

$$\begin{aligned} f_{\vec{X}}(\vec{x}) &= \left(\frac{1}{2\pi}\right)^{\frac{n}{2}} \cdot \frac{1}{\sqrt{\det(\Sigma)}} \cdot \exp\left(-\frac{1}{2}(\vec{x} - \vec{\mu})^{\text{tr}} \cdot \Sigma^{-1} \cdot (\vec{x} - \vec{\mu})\right) \\ &= \prod_{i=1}^n \frac{1}{\sqrt{2\pi}} \cdot \frac{1}{\sigma_i} \cdot \exp\left(-\frac{(x_i - \mu_i)^2}{2\sigma_i^2}\right) \end{aligned}$$

which shows that the joint density factorizes and therefore we can conclude the independence of the components. ■

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Definition (Stopping time - cf. Definition 14.7)

Let  $(\Omega, \mathcal{A})$  be a measurable space and  $(\mathcal{A}_n)_{n \in \mathbb{N}_0}$  a filtration of  $(\Omega, \mathcal{A})$ . A **stopping time** is a function

$$T : \Omega \rightarrow \{0, 1, 2, \dots\} \cup \{\infty\}$$

such that

$$\{T = n\} \in \mathcal{A}_n$$

for all  $n \in \mathbb{N}_0$ .

## Theorem (The optional stopping theorem - cf. Theorem 14.9)

Let  $(\Omega, \mathcal{A}, P)$  be a probability space and  $(\mathcal{A}_n)_{n \in \mathbb{N}_0}$  a filtration of  $(\Omega, \mathcal{A})$ :

- 1 If  $(M_n)_{n \in \mathbb{N}_0}$  is a martingale with respect to  $(\mathcal{A}_n)_{n \in \mathbb{N}_0}$  and  $T$  a stopping time, then the stopped martingale

$$(M_{T \wedge n})_{n \in \mathbb{N}_0}$$

is again a martingale with respect to  $(\mathcal{A}_n)_{n \in \mathbb{N}_0}$ .

- 2 If in addition  $P(T < \infty) = 1$  and  $(M_{T \wedge n})_{n \in \mathbb{N}_0}$  is uniformly integrable, then

$$E[M_T] = E[M_0].$$

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# Sufficient conditions for uniform integrability

Let  $I$  be some index set and  $(X_i)_{i \in I}$  a family of random variables. Each of the following conditions is sufficient for the uniform integrability of  $(X_i)_{i \in I}$ :

- 1 The index set  $I$  is finite and we have  $X_i \in \mathcal{L}^1 \forall i \in I$ .
- 2 There exists a uniform integrable family of random variables  $(Y_i)_{i \in I}$  such that  $|X_i| \leq |Y_i| \forall i \in I$ .
- 3 For some  $p > 1$  we have  $X_i \in \mathcal{L}^p \forall i \in I$ .



Allan Gut.

*An Intermediate Course in Probability.*

Springer, 1995.

(*Chapter 5: The Multivariate Normal Distribution*)



Achim Klenke.

*Probability Theory.*

Springer, 2008.

(*Chapter 6: Convergence Theorems (includes a lot of information about uniform integrability)*)